



¿Cómo corregir la heterocedasticidad y autocorrelación de residuales en modelos de ahusamiento y crecimiento en altura?

How to correct the heteroscedasticity and autocorrelation of residuals in taper and height growth models?

Gerónimo Quiñonez-Barraza^{1*}, Guadalupe Geraldine García-Espinoza² y Oscar Alberto Aguirre-Calderón³

Resumen:

En la modelación del ahusamiento y del crecimiento en altura dominante con datos de series de tiempo, es muy común la presencia de heterocedasticidad y autocorrelación de los errores. Funciones de varianza (varFunc) y estructuras de correlación (corStruct) para corregir la heterocedasticidad y modelar dependencia de los errores, respectivamente. Estas fueron combinadas y evaluadas en ecuaciones de ahusamiento y crecimiento en altura de *Pinus teocote* en Durango, México. La base de datos se obtuvo de 51 análisis troncales con 768 observaciones de ahusamiento y 634 de altura. Las varFunc utilizadas fueron: 1) función de potencia (varPower); 2) función exponencial (varExp); 3) función constante y de potencia (varConstPower); y 4) función combinada de potencia y exponencial (varComb). Las corStruct incluyeron: simetría compuesta (corCompSymm), autorregresiva de orden 1 (corAR1), autorregresiva continua (corCAR1), autorregresiva de media móvil (corARMA2-0), corARMA1-1, corARMA2-1, corARMA2-2, corARMA3-1 y corARMA3-2. Las ecuaciones se ajustaron por mínimos cuadrados generalizados no lineales; y se evaluaron con un sistema de calificación con los estadísticos de ajuste: RMSE, R^2 , AIC, BIC, LogLik, CV y sesgo promedio. Con base en la calificación, las mejores combinaciones para el ahusamiento y crecimiento en altura fueron 1-9, 2-5, 3-8 y 4-6 y 1-6, 2-9, 3-7 y 4-4, respectivamente. En el ahusamiento solo la combinación 2-5 fue homocedástica con residuales independientes al igual que las ecuaciones de altura seleccionadas y las varFunc y corStruct presentaron influencia en la trayectoria de las curvas de índice de sitio construidas.

Palabras clave: Ahusamiento, altura dominante, estructuras de correlación, funciones de varianza, *Pinus teocote* Schiede ex Schltdl. & Cham., residuales.

Abstract:

In modeling of taper functions and dominant height growth with time series data, the presence of heteroscedasticity and autocorrelation in residuals is common. Variance Functions (varFunc) and correlation structures (corStruct) were used to correct heteroscedasticity and autocorrelation; both were combined and evaluated through taper and height growth equations for *Pinus teocote* in Durango, Mexico. A dataset of 51 stems analysis with 768 taper observations and 634 height growth observations was used. The varFuncs applied were: 1) power function (varPower); 2) exponential function (varExp); 3) constant plus power function (varConstPower); and 4) a combination of power and exponential functions (varComb). The corStructs were: compound symmetry (corCompSymm), autoregressive of order 1 (corAR1), continuous-time autoregressive of order 1 (corCAR1), autoregressive-moving average (corARMA2-0), corARMA1-1, corARMA2-1, corARMA2-2, corARMA3-1 and corARMA3-2. To fit the equations, the generalized nonlinear least squares method was used and evaluated with a rating system through: RMSE, R^2 , AIC, BIC, LogLik, VC and average bias. According to the rating system, the best combinations for taper and height growth equations were 1-9, 2-5, 3-8 and 4-6 and 1-6, 2-9, 3-7 and 4-4, respectively. In the taper equation, only the combination 2-5 was homoscedastic with independent residuals, and the selected height growth equations were homoscedastic with independent residuals; the varFunc and corStruct had influence on the trajectories of site index curves.

Key words: Taper, dominant height, correlation structures, variance functions, *Pinus teocote* Schiede ex Schltdl. & Cham., residuals.

Fecha de recepción/Reception date: 10 de diciembre de 2017

Fecha de aceptación/Acceptance date: 30 de julio de 2018

¹Campo Experimental Valle del Guadiana, INIFAP. México. correo-e: quinonez.geronimo@inifap.gob.mx

²Posgrado en Ciencias Forestales, Universidad Autónoma de Nuevo León. México.

³Facultad de Ciencias Forestales, Universidad Autónoma de Nuevo León. México.

Introduction

The planning, implementation and monitoring of sustainable forest management require research to support decision-making and the assessment of the established goals. The estimation of the timber stock and productivity of the stands is a main objective in forest management systems; therefore, it is essential to know the growth of commercial species (Aguirre-Calderón, 2015; Salas *et al.*, 2016).

Research on the estimation of volume, growth and increment is a key tool for understanding the dynamics of ecosystems involved in forest management; therefore, these approaches continue to be necessary for the planning and implementation of forest activities. Taper and dominant height growth equations have been widely explored topics (Castillo *et al.*, 2013; Paulo *et al.*, 2015; Rodríguez-Carrillo *et al.*, 2015; Krisnawati, 2016; Corral-Rivas *et al.*, 2017; Fierros-Mateo *et al.*, 2017; Tamarit *et al.*, 2017) because they are biometric tools that allow characterizing the profile of the trees, as well as estimating the total volume and the commercial volume, and the forest productivity, based on the site index approach (Crecente-Campo *et al.*, 2013; Quiñonez-Barraza *et al.*, 2015; Özçelik and Crecente-Campo, 2016).

Since the datasets used for fitting taper and height growth equations are time series obtained from measured variables on the same tree and resulting from taper and tree stem analyses, it is reasonable to assume that the observations in each tree are correlated and, also the residuals of the adjusted equations (Arias-Rodil *et al.*, 2015; Quiñonez-Barraza *et al.*, 2015; Corral-Rivas *et al.*, 2017); furthermore, their predictions might show a variation in the levels of the predictor variables, which is known as heteroscedasticity (Gujarati and Porter, 2011).

The term autocorrelation refers to the correlation between the residuals of a regression model when series of observations arranged in time are used, *e.g.* time-series data, or in space, and with cross-sectional data. The linear and nonlinear regression models are based on the theoretical assumption that the residues have the same variance and are, therefore, homoscedastic. The presence of

autocorrelation and heteroscedasticity leads to estimations of non-minimum variance parameters and to somewhat unreliable prediction intervals, especially in modelling taper and volume equations (Fortin *et al.*, 2007; Xu *et al.*, 2014; Tang *et al.*, 2016). Consequently, the usual t or F tests are not valid. Therefore, the use of generalized least squares (GNLS) approach with variance functions and correlation structures is an alternative to generate the best unbiased linear parameter estimates (Gujarati and Porter, 2011).

In studies of taper and dominant height growth models, correlation structures and power functions to correct the autocorrelation and heteroscedasticity of the residuals, respectively, are frequently used (Quiñonez-Barraza *et al.*, 2014; Quiñonez-Barraza *et al.*, 2015; Sharma *et al.*, 2015; Özçelik and Crecente-Campo, 2016; Corral-Rivas *et al.*, 2017; Tamarit *et al.*, 2017); continuous-time autoregressive structures of errors with one, two or three power functions with known variance are highlighted.

Because it is important to improve the predictive capacity and the interpretation of the statistic properties in equations fitting, the objective of this study was to evaluate the combination of variance functions with correlation structures, in order to model the heteroscedasticity and the errors dependence in taper and dominant height growth equations of *Pinus teocote* Schiede ex Schltdl. & Cham. in *Durango*, Mexico.

Materials and Methods

Study area and description of experimental data

The information from stem analysis of 51 *Pinus teocote* trees collected in mixed stands of the Forest Management Unit (Umafor) 1005 *Santiago Papasquiaro y Anexos*, in northeastern *Durango*, Mexico, was used. The study area is located on the *Sierra Madre Occidental*, between 24°48'16.98" and 25°13'47.25" N, and 105°53'9.81" and 106°12'52.58" W. The forest polygon was *San Diego de Tezains*

ejido, with a total area of 61 098.25 ha, of which 26 636.09 ha are programmed for timber production with forest management (Quiñonez-Barraza et al., 2014). The predominant climates are temperate warm humid and temperate subhumid, with a mean annual precipitation of 1 375 mm. The mean annual temperatures ranges from 8 °C in the highest areas to 24 °C in the lower areas, where the mean altitude is around 600 m (García, 1981; Quiñonez-Barraza et al., 2015).

The data was taken using a totally random design for the stands of the timber production area, and was considered a normal distribution for the diameter categories. The trees were felled and divided into sections in order to register the growth in dominant height and taper by relative heights. The first measurement corresponded to the stump height; subsequently, to lengths of 0.30 m, 0.60 m and 1.3 m; furthermore, growth slices were extracted every 2 m (Quiñonez-Barraza et al., 2015), whose diameters, heights and number of rings were recorded. In whole dataset, 768 diameter-height (taper) and 685 height-age combinations were used. Table 1 shows the descriptive statistics of the analyzed variables.

Table 1. Descriptive statistics of the analyzed variables in order to fit the taper and dominant height growth of *Pinus teocote* Schiede ex Schltdl. & Cham.

Statistic	D	d	H	h	hs	ds	Ac	At
Minimum	13.00	0.00	7.85	0.00	0.10	19.00	3.00	34.00
Maximum	49.00	62.00	26.60	26.60	0.35	62.00	172.00	172.00
Mean	26.60	19.78	15.80	7.16	0.19	36.66	37.16	79.33
SD	9.45	11.79	4.34	6.15	0.05	10.74	27.83	28.96

D = Normal diameter (cm); d = Diameter at commercial height h (cm); H = Total height (m); h = Commercial height (m); hs = Stump height (m); ds = Diameter at stump height (cm); Ac = Age at commercial height h (years); At = Total age (years); SD = Standard deviation from the mean.

Fitting models

The taper was modeled with the segmented equation developed by Fang *et al.* (2000), which has been used in several studies in order to generate compatible taper and merchantable volume systems in species of interest (Quiñonez-Barraza *et al.*, 2014; Uranga-Valencia *et al.*, 2015; Özçelik and Crecente-Campo, 2016; Corral-Rivas *et al.*, 2017; Tamarit *et al.*, 2017). The segmented taper equation is as follows:

$$d_{ij} = c_1 \left(H_i^{(K-\beta_1)/\beta_1} \left(1 - \frac{h_{ij}}{H_i} \right)^{(K-R)/R} A_1^{(I_1+I_2)} A_2^{I_2} \right)^{0.5} + \varepsilon_{ij} \quad (1)$$

Where:

$$c_1 = \left(\alpha_0 D_i^{\alpha_1} H_i^{(\alpha_2-K)/\beta_1} / [\beta_1(t_0 - t_1) + \beta_2(t_1 - A_1 t_2) + \beta_3 A_1 t_2] \right)^{0.5}$$

$$t_0 = (1 - \rho_0)^{K/\beta_1}$$

$$\rho_0 = h_{bi}/H_i$$

$$t_1 = (1 - \rho_1)^{K/\beta_1}$$

$$t_2 = (1 - \rho_2)^{K/\beta_2}$$

$$A_1 = (1 - \rho_1)^{(\beta_2-\beta_1)K/\beta_1\beta_2}$$

$$A_2 = (1 - \rho_2)^{(\beta_3-\beta_2)K/\beta_2\beta_3}$$

$$R = \beta_1^{1-(I_1+I_2)} \beta_2^{I_1} \beta_3^{I_2}$$

$$I_1 = \begin{cases} 1 & \text{if } \rho_1 \leq z \leq \rho_2 \\ 0 & \text{otherwise} \end{cases}; I_2 = \begin{cases} 1 & \text{if } \rho_2 < z < 1 \\ 0 & \text{otherwise} \end{cases}$$

$$\rho_1 = h_{ij1}/H_i$$

$$\rho_2 = h_{ij2}/H_i$$

$$d_{ij} = \text{Diameter } j \text{ of the tree } i \text{ at the commercial height } h_{ij} \text{ (cm)}$$

$$H_i = \text{Total height of the tree } i \text{ (m)}$$

$$h_{ij} = \text{Commercial height } j \text{ of the tree } i \text{ (m)}$$

D_i = Normal diameter of the tree i (cm)

h_{bi} = Stump height of the tree i (m)

α_i = Total volume parameters ($i = 1, 2, 3$)

β_i = Taper parameters ($i = 1, 2, 3$)

ε_{ij} = Residual of diameter j in the tree i

The dominant height as an intrinsic site index equation was modeled using the dynamic equation in the generalized algebraic difference approach (GADA), derived by Quiñonez-Barraza *et al.* (2015) and expressed as equation 2.

$$h_{2ij} = e^{\left[\beta_1 + \beta_2 \left(\frac{\ln(h_{1ij}) - \beta_1}{\ln(1 - e^{-\beta_3 t_{1ij}}) + \beta_2} \right) \right]} (1 - e^{-\beta_3 t_{2ij}})^{\left[\frac{\ln(h_{1ij}) - \beta_1}{\ln(1 - e^{-\beta_3 t_{1ij}}) + \beta_2} \right]} + \varepsilon_{ij} \quad (2)$$

Where:

h_{2ij} = Height j of the tree i at age j of the tree i in status 2 (m)

h_{1ij} = Height j the tree i at age j of the tree i in status 1 (m)

t_{1ij} = Age j of the tree i in status 1 (years)

t_{2ij} = Age j of the tree i in status 2 (years)

β_i = Parameters to be estimated ($i = 1, 2, 3$)

e = Exponential function

$\ln$ = Natural logarithm

ε_{ij} = Residual of height j in the tree i

Autocorrelation and heteroscedasticity

Combinations of variance functions with correlation structures were used in taper equation (Eq. 1) and dominant height equation (Eq. 2). The varFunc and corStruct were determined according to Pinheiro and Bates (2000).

Variance functions. The variance functions were as follow: 1) power function (varPower); 2) exponential function (varExp); 3) constant power function (varConstPower), and 4) combination of power-exponential functions (varComb). The variance functions were used in order to model the variability between the measurements of each tree i with the merchantable height covariables (h_{ij}) for the taper equation, and the dominant height (h_{1ij}) at state t_{1ij} for the height growth equation. The general structure of the power functions for modeling the heteroscedasticity considers two arguments for most varFuncs: the parameter value and shape. The first specifies the value of the variance parameter (δ), and the second specifies the variance covariable (v), which in this case may be considered a stratified variable for each tree; *i.e.* a variance parameter can be estimated for each dataset of the several trees (Pinheiro and Bates, 2000).

The variance model (Var or σ^2) for varPower is represented by equation 3 and corresponds to the variance function (g) in equation 4. Covariable h_{ij} was utilized for the taper equation (Eq. 1), while h_{1ij} was used for the dominant height growth equation (Eq. 2).

$$Var(\varepsilon_{ij}) = \sigma^2 |v_{ij}|^{2\delta} \quad (3)$$

$$g(v_{ij}, \delta) = |v_{ij}|^\delta \quad (4)$$

The varExp variance model of is represented by equation 5, and the corresponding function, by equation 6, for the same covariables previously defined for taper and dominant height growth equation. The varConstPower variance model is defined in equation 7, and the variance function, in equation 8. The varComb variance model

(varExp and varPower) is defined in equation 9, with the respective function expressed in equation 10 (Pinheiro and Bates, 2000). In all cases, the same covariables, previously defined, were used. For the varConstPower, δ_1 represents the constant parameter, and δ_2 , the corresponding power parameter; δ_1 corresponds to the varExp parameter, while δ_2 corresponds to varPower.

$$Var(\varepsilon_{ij}) = \sigma^2 e^{(2\delta v_{ij})} \quad (5)$$

$$g(v_{ij}, \delta) = e^{(\delta v_{ij})} \quad (6)$$

$$Var(\varepsilon_{ij}) = \sigma^2 \left(\delta_1 + |v_{ij}|^{\delta_2} \right)^2 \quad (7)$$

$$g(v_{ij}, \delta) = \delta_1 + |v_{ij}|^{\delta_2} \quad (8)$$

$$Var1(\varepsilon_{ij}) \times Var2(\varepsilon_{ij}) = \sigma^2 e^{(2\delta_1 v_{ij})} \times \sigma^2 |v_{ij}|^{2\delta_2} \quad (9)$$

$$g1(v_{ij}, \delta_1) \times g2(v_{ij}, \delta_2) = e^{(\delta_1 v_{ij})} \times |v_{ij}|^{\delta_2} \quad (10)$$

The correlation structures were: 1) compound symmetry (corCompSymm); 2) autoregressive of order 1 (corAR1); 3) continuous-time autoregressive of order 1 (corCAR1); 4) autoregressive-moving average 2, 0 (corARMA2-0); 5) autoregressive-moving average 1, 1 (corARMA1-1); 6) autoregressive-moving average 2, 1 (corARMA2-1); 7) autoregressive-moving average 2, 2 (corARMA2-2); 8) autoregressive-moving average 3, 1 (corARMA3-1), and 9) autoregressive-moving average 3, 2 (corARMA3-2). The correlation structures were used to model the dependence between residuals of each tree, with time-series data (Pinheiro and

Bates, 2000). This study modeled the dependence between the diameter and height measurements in the same tree for the aim of achieving independence in the residuals of taper equation (Eq. 1) and dominant height growth equation (Eq. 2). The general structure of correlation between groups for a single grouping level is expressed as equation 11 (Pinheiro and Bates, 2000).

$$\text{cor}(\varepsilon_{ij}, \varepsilon_{i1j-1}) = f[d(p_{i1j}, p_{i1j-1}), \rho] \quad (11)$$

Where:

ε_{ij} = Residual of the diameter or height j of the tree i

ε_{i1j-1} = Residual of the diameter or delayed height growth $j - 1$ in the tree i

f = Correlation function $f(\cdot)$, taking up values between 1 and -1

d = Distance between position vectors p_{i1j} and p_{i1j-1} , which, for the taper equation, were the differences between the commercial height h_{ij} and h_{ij-1} ; while, for the dominant height growth equation, they were the heights defined in two positions: h_{1ij} and h_{1ij-1}

ρ = Correlation parameters vector

In the correlation structure `corCompSymm`, an equal correlation is assumed for all errors of the same group within the same tree; the correlation model is given by equation 12. The `corAR1` model is represented in equation 13, while the `carCAR1` model is expressed by equation 14. The autoregressive-moving average (`corARMA2-0`) is defined by equation 15, and `corARMA1-1`, by equation 16, which is the generalization for `corARMA` p , q structures; i.e. `corARMA2-0`, `corARMA1-1`, `corARMA2-1`, `corARMA2-2`, `corARMA3-1`, and `corARMA3-2` are represented by the changes in p values for the autoregressive structure (AR), and q , for the moving average (MA) structure. The new error term (ε_{ij}) defines the dependent residuals in the regression model; in the moving average (MA) structure, the residual is determined by a_{ij} .

$$\varepsilon_{ij} = \rho_1 [d(p_{i1j}, p_{i1j-1})] + \epsilon_{ij} \quad (12)$$

$$\varepsilon_{ij} = \rho_1 \varepsilon_{ij-1} + \epsilon_{ij} \quad (13)$$

$$\varepsilon_{ij} = \rho_1 [d(p_{i1j}, p_{i1j-1})] \varepsilon_{ij-1} + \epsilon_{ij} \quad (14)$$

$$\varepsilon_{ij} = \sum_{k=1}^{q=2} \rho_k \varepsilon_{ij-k} + \epsilon_{ij} \quad (15)$$

$$\varepsilon_{ij} = \sum_{k=1}^{p=1} \rho_k \varepsilon_{ij-k} + \sum_{k=1}^{q=1} \theta_1 a_{ij-k} + a_{ij} + \epsilon_{ij} \quad (16)$$

Fitting equations and goodness-of-fit statistics

The equations were fitted with generalized nonlinear least squares (GNLS) of the statistical package NLME (nonlinear mixed effects models) (Pinheiro *et al.*, 2015), in the R environment (R Core Team, 2017). In order to assess the fit of taper and dominant height growth equations, combinations of the variance functions with correlation structures like those studied by Pinheiro and Bates (2000) were used. The goodness-of-fit was evaluated using the following six statistics: root mean square error (RMSE), adjusted coefficient of determination (R^2_a), Akaike Information Criterion (AIC), Bayesian Information Criterion (BIC), Log-Likelihood (logLik), coefficient of variation (CV) and absolute mean bias (Bias). A rating system was generated with these statistics in order to select the best varFunc and corStruct combinations. Each statistic was assigned a value from 1 to 9; 1 corresponds to the combination with the best statistic, and 9, to the one with the worst statistic (Sakici *et al.*, 2008; Tamarit *et al.*, 2014). The Durbin-Watson (Dw) statistic (Durbin and Watson, 1971) was used to evaluate the correction of the autocorrelation, with a robust modification

(DwM), such as the average Dw between groups, since errors are regarded as dependent on the measurements of each tree, but not on the general dataset.

The modified statistic is shown in equation 17. The homogeneity of variances was tested using Bartlett's test (Bartlett, 1954; Layard, 1973; Hidding *et al.*, 2013); the variation between the groups and a significance value of 1 % ($\alpha = 1\%$) were considered for taper and dominant height growth equations. The Assumption of homogeneity of variance (null hypothesis, H0) is expressed in equation 18; therefore, higher values than 0.01 of probability value assume variance homogeneity, and, conversely (alternate hypothesis, Ha), lower values than 0.01 imply that at least two variances are different.

$$DwM = \frac{\frac{\sum_{j=2}^J (\varepsilon_{ij} - \varepsilon_{ij-1})^2}{\sum_{j=2}^J \varepsilon_{ij}^2}}{N} \quad (17)$$

Where:

DwM = Modified Durbin-Watson statistic

ε_{ij} = Diameter or height j residual for the tree i

J = Number of observations in each tree i

N = Number of trees i in the database

$$X^2 = \frac{(N-k) \ln(s_p^2) - \sum_{i=1}^k (n_i - 1) \ln(s_i^2)}{1 + \frac{1}{3(k-1)} \left(\sum_{i=1}^k \left(\frac{1}{n_i - 1} \right) - \frac{1}{N-k} \right)} \quad (18)$$

Where:

K = Samples of size n_i

S_i^2 = Sample variance

$$N = \sum_{i=1}^k n_i$$

$$S_p^2 = \frac{1}{N-k} \sum_i (n_i - 1) S_i^2 \text{ (estimation of the combined variance)}$$

Results and Discussion

The combinations of variance functions and correlation structures generated 36 models for taper, and 36 for height growth, based on equations 1 and 2, respectively. The fit statistics and the ranking score (RS) showed the goodness-of-fit of the equations (Table 2, for taper, and Table 3, for height growth). Also the DwM statistics and probability of Bartlett's test of homogeneity of variances (P -value).

The ranking score exhibited the combinations of variance functions with correlation structures by hierarchical order (Sakici *et al.*, 2008). The lowest RS value was the statistically best combination, and the highest RS corresponded to the worst combination, based on the sum of the ranks for each fitting statistic (Tamarit *et al.*, 2014).

In the case of taper, the varPower function combined with corARMA3-2 performed best, while the worst combination was with corCAR1. In all cases of correlation structures combined with varPower, values ranges from 1.634 to 1.862 were displayed for the DwM statistic. However, for the test of homogeneity of variances, all the combinations were heteroscedastic. In the variance function varExp, all the combinations with correlation structures, except for corCompSymm, were homoscedastic ($P > 0.01$). And the corARMA1-1 structure performed best according to the RS, with a DwM value of 1.553.

The varConstPower and varComb functions combined with the correlation structures have consistent statistics and independent residuals (DwM values from 1.722 to 1.973); however, they exhibit heteroscedasticity in Bartlett's test. Structures corARMA3-1 and corARMA2-1 were the best for varConstPower and varComb, respectively.

In combinations of variance functions and correlation structures, the height growth generated DwM statistics of approximately 1.12. However, for most equations, they exhibited homogeneous variances. The varPower function is likewise statistically combined with the corARMA2-1, corARMA2-2 and corARMA3-2 structures, since the RS was 22 for all three cases. The corCAR1 structure had the lowest fitting, with unequal variances. Table 3 shows the behavior of the variance functions varExp, varConstPower and varComb with the combinations of correlation structures.

Table 2. Adjustment statistics of the taper equations for the combinations of variance functions with correlation structures.

Combination	RMSE	R ² a	AIC	BIC	logLik	CV	Bias	RS	DwM	P-Value
H1-A9	1.900	0.974	2474	2544	-1222	9.391	0.313	20	1.821	<0.001
H1-A2	1.945	0.972	2475	2526	-1227	9.510	0.446	28	1.770	<0.001
H1-A7	1.946	0.972	2475	2540	-1223	9.490	0.453	30	1.793	<0.001
H1-A1	1.839	0.975	2799	2850	-1389	9.186	0.203	31	1.634	<0.001
H1-A6	1.949	0.972	2474	2534	-1224	9.486	0.473	33	1.770	<0.001
H1-A8	2.021	0.970	2444	2509	-1208	9.700	0.582	35	1.822	<0.001
H1-A5	1.949	0.972	2476	2532	-1226	9.517	0.452	38	1.794	<0.001
H1-A4	1.951	0.972	2476	2532	-1226	9.514	0.464	40	1.770	<0.001
H1-A3	2.195	0.965	2501	2552	-1239	10.577	0.621	60	1.868	<0.001
H2-A5	1.751	0.978	2585	2641	-1280	8.792	0.003	19	1.553	0.137
H2-A4	1.751	0.978	2585	2640	-1280	8.793	0.003	20	1.553	0.137
H2-A7	1.759	0.977	2585	2650	-1278	8.825	-0.022	25	1.545	0.318
H2-A6	1.753	0.977	2587	2648	-1281	8.800	-0.014	28	1.549	0.217
H2-A8	1.759	0.977	2586	2651	-1279	8.822	-0.019	28	1.544	0.142
H2-A2	1.869	0.974	2687	2738	-1332	9.126	-0.438	45	1.322	0.211
H2-A9	2.058	0.969	2678	2748	-1324	9.732	-0.677	47	1.104	0.322

H2-A1	1.769	0.977	2867	2918	-1423	8.889	-0.032	48	1.634	<0.001
H2-A3	1.867	0.974	2809	2860	-1393	9.375	0.081	55	1.380	0.095
H3-A8	1.935	0.973	2476	2545	-1223	9.476	0.409	16	1.789	<0.001
H3-A2	1.946	0.972	2477	2533	-1227	9.510	0.446	29	1.770	<0.001
H3-A7	1.938	0.973	2479	2548	-1224	9.478	0.419	29	1.786	<0.001
H3-A1	1.801	0.976	2772	2828	-1374	8.993	0.189	31	1.722	<0.001
H3-A6	1.951	0.972	2476	2541	-1224	9.486	0.473	32	1.770	<0.001
H3-A5	1.950	0.972	2478	2539	-1226	9.517	0.452	37	1.794	<0.001
H3-A4	1.952	0.972	2478	2538	-1226	9.514	0.464	38	1.791	<0.001
H3-A9	1.955	0.972	2478	2553	-1223	9.500	0.468	43	1.769	<0.001
H3-A3	2.197	0.965	2503	2559	-1239	10.577	0.621	60	1.868	<0.001
H4-A6	1.972	0.972	2401	2466	-1186	9.674	0.408	24	1.953	<0.001
H4-A9	1.964	0.972	2401	2476	-1185	9.705	0.319	24	1.973	<0.001
H4-A8	1.974	0.972	2403	2473	-1186	9.675	0.408	30	1.953	<0.001
H4-A7	1.974	0.972	2403	2473	-1186	9.675	0.408	31	1.953	<0.001
H4-A4	2.074	0.969	2366	2426	-1170	9.845	0.668	36	1.910	<0.001
H4-A5	1.977	0.971	2404	2464	-1189	9.668	0.440	36	1.940	<0.001
H4-A1	1.959	0.972	2710	2761	-1344	9.517	0.492	37	1.931	<0.001
H4-A2	1.977	0.971	2403	2458	-1189	9.686	0.433	38	1.941	<0.001
H4-A3	2.131	0.967	2492	2548	-1234	10.342	0.543	59	1.902	<0.001

H1 = varPower; H2 = varExp; H3 = varConstPower; H4 = varComb;
A1 = corCompSymm; A2 = corAR1; A3 = corCAR1; A4 = corARMA2-0; A5 = corARMA1-1;
A6 = corARMA2-1; A7 = corARMA2-2; A8 = corARMA3-1; A9 = corARMA3-2; RMSE = Root
mean-square error; R^2_a = Adjusted coefficient of determination; AIC = Akaike
Information Criterion; BIC = Bayesian information criterion; logLik = Log-
Likelihood; CV = Coefficient of variation; Bias = Absolute mean bias; RS = Total
ranking score to the rating system; P-Value = probability of Bartlett's test.

Table 3. Adjustment statistics for height growth equations for the combinations of variance functions with correlation structures.

Combination	RMSE	R ² a	AIC	BIC	logLik	CV	Bias	RS	DwM	P-Value
H1-A6	0.793	0.981	1013	1048	-498	6.597	0.552	22	1.206	0.020
H1-A7	0.792	0.981	1013	1053	-497	6.601	0.549	22	1.215	0.018
H1-A9	0.792	0.981	1015	1059	-497	6.601	0.548	22	1.218	0.018
H1-A5	0.793	0.981	1011	1042	-498	6.597	0.552	25	1.206	0.019
H1-A8	0.794	0.981	1014	1054	-498	6.596	0.552	31	1.204	0.020
H1-A1	0.760	0.983	1148	1175	-567	6.712	0.494	39	1.438	<0.001
H1-A4	0.796	0.981	1023	1054	-504	6.607	0.555	49	1.182	0.018
H1-A2	0.796	0.981	1076	1103	-532	6.620	0.555	50	1.161	0.034
H1-A3	0.796	0.981	1076	1103	-532	6.620	0.555	55	1.161	<0.001
H2-A9	0.798	0.981	982	1026	-480	6.587	0.558	25	1.176	0.018
H2-A7	0.798	0.981	980	1020	-481	6.587	0.559	28	1.172	0.018
H2-A5	0.798	0.981	977	1008	-481	6.586	0.561	30	1.167	0.020
H2-A6	0.799	0.981	979	1014	-481	6.586	0.561	30	1.168	0.001
H2-A8	0.800	0.981	980	1020	-481	6.585	0.561	35	1.176	0.018
H2-A2	0.798	0.981	1054	1081	-520	6.617	0.559	36	1.148	0.040
H2-A1	0.756	0.983	1126	1153	-556	6.626	0.496	38	1.398	<0.001
H2-A3	0.798	0.981	1054	1081	-520	6.617	0.559	39	1.148	<0.001
H2-A4	0.824	0.980	998	1030	-492	6.642	0.592	54	1.150	0.021
H3-A7	0.793	0.981	1015	1060	-497	6.601	0.549	22	1.215	0.018
H3-A9	0.793	0.981	1017	1066	-497	6.601	0.548	22	1.218	0.018
H3-A5	0.794	0.981	1013	1048	-498	6.597	0.552	23	1.206	0.019
H3-A6	0.794	0.981	1015	1055	-498	6.597	0.552	24	1.206	0.020
H3-A8	0.795	0.981	1016	1061	-498	6.596	0.552	31	1.204	0.020
H3-A1	0.761	0.983	1150	1181	-567	6.712	0.494	39	1.438	<0.001
H3-A2	0.796	0.981	1078	1110	-532	6.620	0.555	49	1.161	0.034
H3-A4	0.797	0.981	1025	1061	-504	6.607	0.555	49	1.182	0.018
H3-A3	0.796	0.981	1078	1110	-532	6.620	0.555	56	1.161	<0.001
H4-A4	0.834	0.979	941	976	-462	6.666	0.602	28	1.058	0.008
H4-A5	0.855	0.978	917	953	-450	6.684	0.630	28	1.058	0.019
H4-A7	0.853	0.978	920	965	-450	6.679	0.625	29	0.972	0.016

H4-A3	0.820	0.980	1014	1046	-500	6.671	0.584	32	0.835	<0.001
H4-A2	0.820	0.980	1014	1046	-500	6.671	0.584	33	0.529	0.050
H4-A6	0.856	0.978	919	959	-450	6.685	0.630	33	0.961	0.019
H4-A9	0.865	0.978	913	962	-445	6.699	0.639	37	0.942	0.025
H4-A8	0.859	0.978	919	964	-449	6.688	0.632	38	0.957	0.026
H4-A1	0.919	0.975	992	1023	-489	6.813	0.705	57	0.417	0.096

H1 = varPower; H2 = varExp; H3 = varConstPower; H4 = varComb;
A1 = corCompSymm; A2 = corAR1; A3 = corCAR1; A4 = corARMA2-0;
A5 = corARMA1-1; A6 = corARMA2-1; A7 = corARMA2-2; A8 = corARMA3-1;
A9 = corARMA3-2; RMSE = Root Mean Square Error; R^2_a = Adjusted coefficient of
determination; AIC = Akaike Information Criterion; BIC = Bayesian Information
Criterion; logLik = Log-Likelihood; CV = Coefficient of variation; Bias =
Absolute mean bias; RS = Total ranking score; *P*-value = Probability of
Bartlett's test.

In all the fitted equations, the estimated parameters were statistically different from zero at a significance level of 5 % ($P < 0.05$). The estimated parameters for the best combinations of each variance function with the correlation structures are summarized in Table 4, for both taper (T) and height growth (HG). In the case of taper, the combinations varExp&corARMA3-2, varExp&corARMA2-2 and varConstPower&corARMA3-1 represented 14 parameter estimates, and the test of homogeneity of variances displayed unequal variances; while the combination varExp&corARMA1-1, with 12 parameter estimates, exhibited constant variances.

The estimated parameters defining the changes of the dendrometric stem shapes of the segmented taper model were found for the change from neiloid into paraboloid, as 4.0 % to 7.2 %; while the change from paraboloid into cone occurs in average at 70 % of the total height in the tree stem profiles. Similar results have been documented by researchers for different *Pinus* species (Uranga-Valencia *et al.*, 2015; Özçelik and Crecente-Campo, 2016; Corral-Rivas *et al.*, 2017; Tamarit *et al.*, 2017), with values range from 4 % to 7 % for the first inflection point, and range from 55 % to 75 % for the corresponding second inflection point. Furthermore,

continuous-time autoregressive structures of orders 1 and 2 and power functions were used, and a known variance was assumed. Nevertheless, the studies do not include the corresponding test of homogeneity of variances.

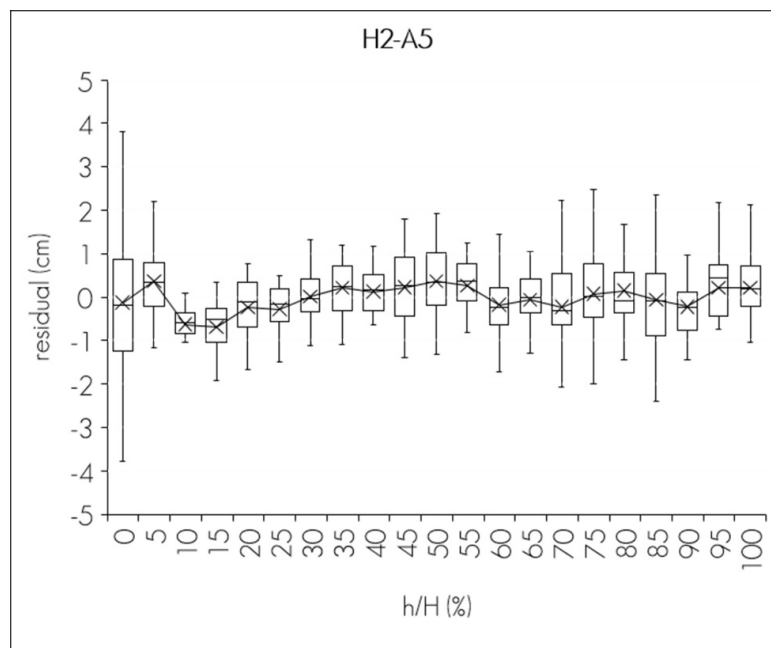
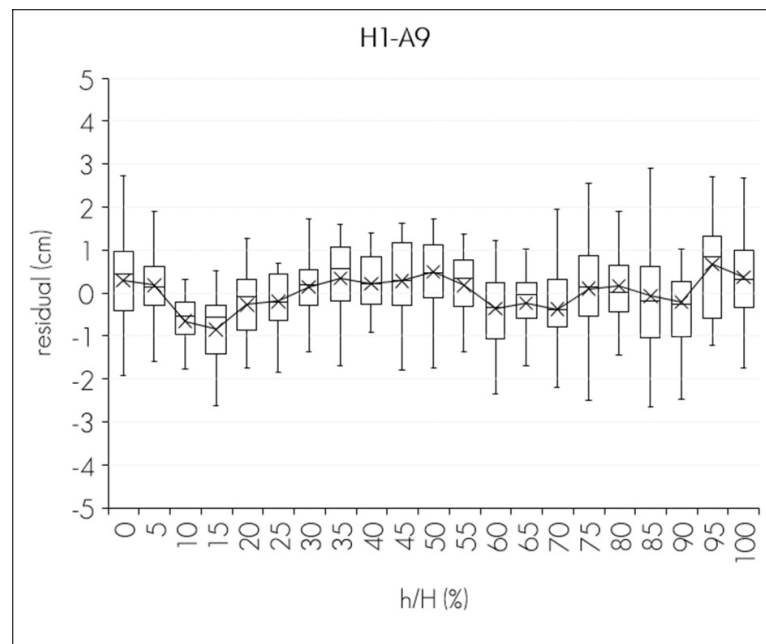
The tendency of the residuals by relative (h/H ; %) of the four best combinations of variance functions with correlation structures for taper are shown as box and whisker plots (Figure 1). Although the tendencies are very similar, only the varExp&corARMA1-1 (H2-A5) combination exhibited a constant variance, and the dispersion of the extreme values are shown closer to the values of the corresponding quantiles.

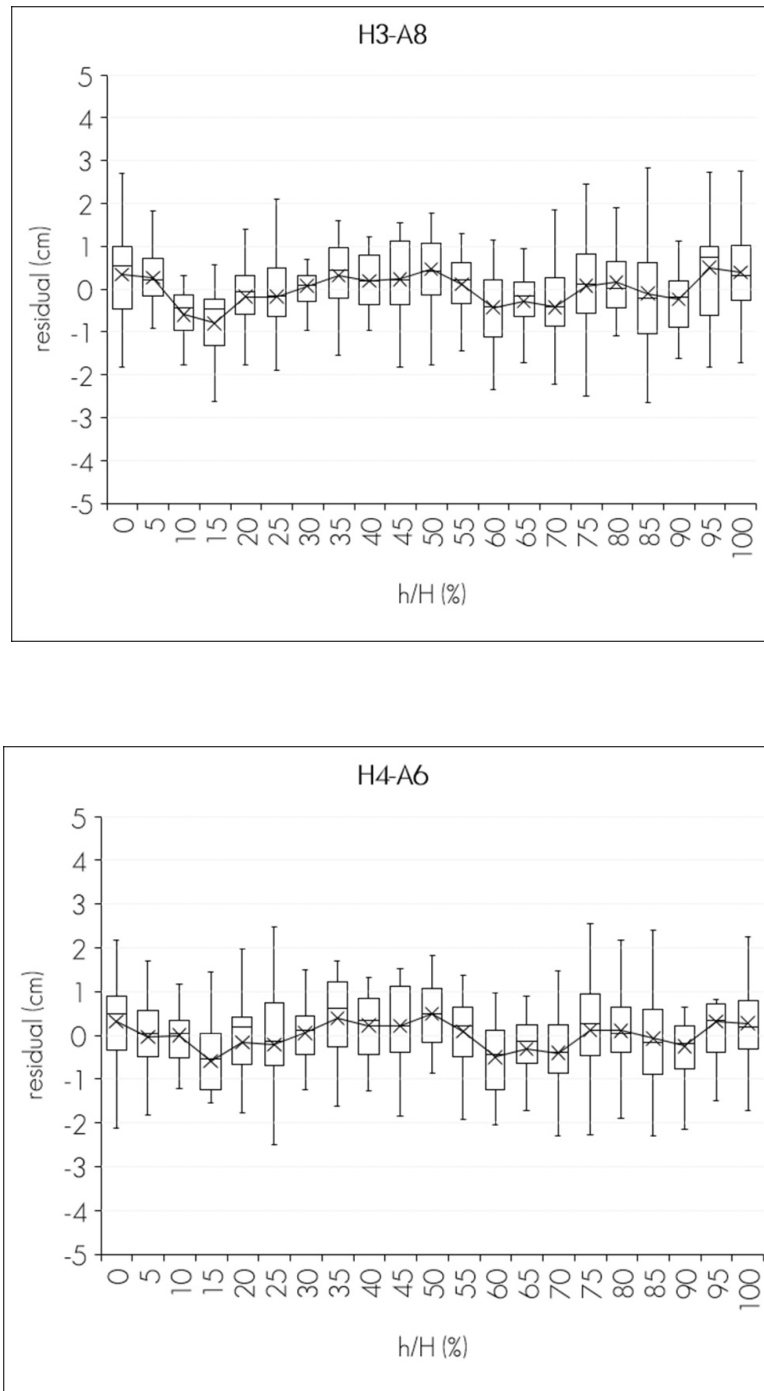
The combinations in the height growth equations with the largest number of parameter estimates (9) were varExp&corARMA3-2 and varConstPower&corARMA2-2; the varExp&corARMA2-1 and varComb&corARMA2-0 combinations had seven estimated parameters.

The use of the equations with multiple parameters can cause an overparameterization, and the predictions resulting from them could not be the most efficient (Gregoire and Schabenberger, 1996). In the four cases, the variances were constant, which guarantees that the parameters are unbiased and efficient and have a minimum variance (Gujarati and Porter, 2011; Tang *et al.*, 2016).

The selected combinations of the height growth equations generated a dispersion of residuals close to zero line, which appear in the form of boxes and whisker plots by relative height (h/H , %) in Figure 2. Although there is an overestimation from 1 % to 10 % relative height classes, and an underestimation from 10 % to 25 % relative height classes, the variances were constant for all four cases, according to Bartlett's test of homogeneity of variances (Bartlett, 1954; Hidding *et al.*, 2013).

As for the residuals, although the DwM test generated values of approximately 1.2, they were considered to be independent within the measurements of height and age carried out in the same tree (Crecente-Campo *et al.*, 2013; Quiñonez-Barraza *et al.*, 2015), and therefore the estimated parameters are unbiased, consistent, and efficient (Williams *et al.*, 2013).

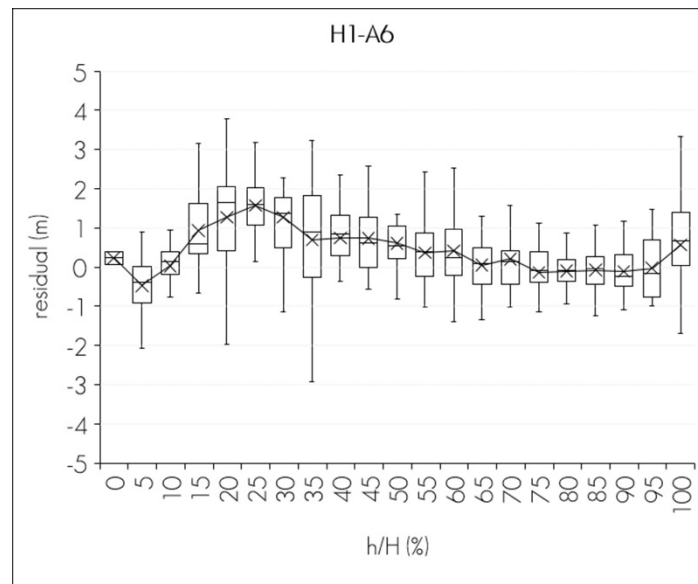


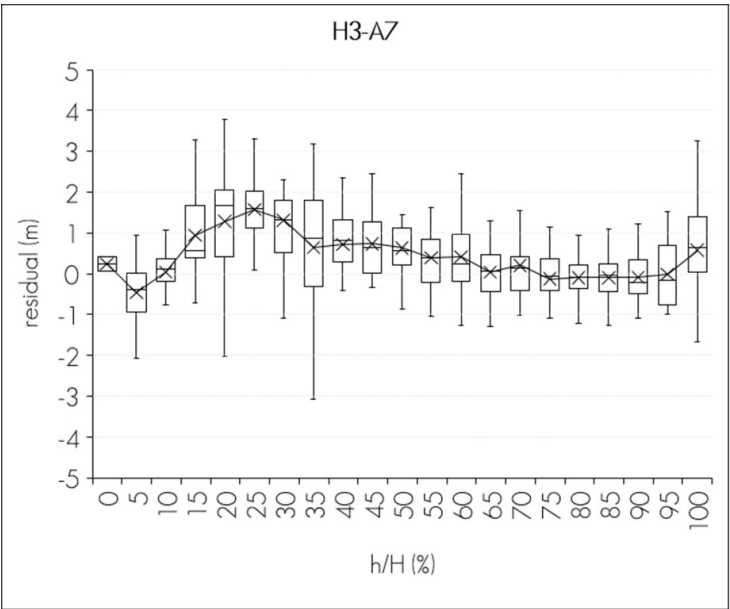
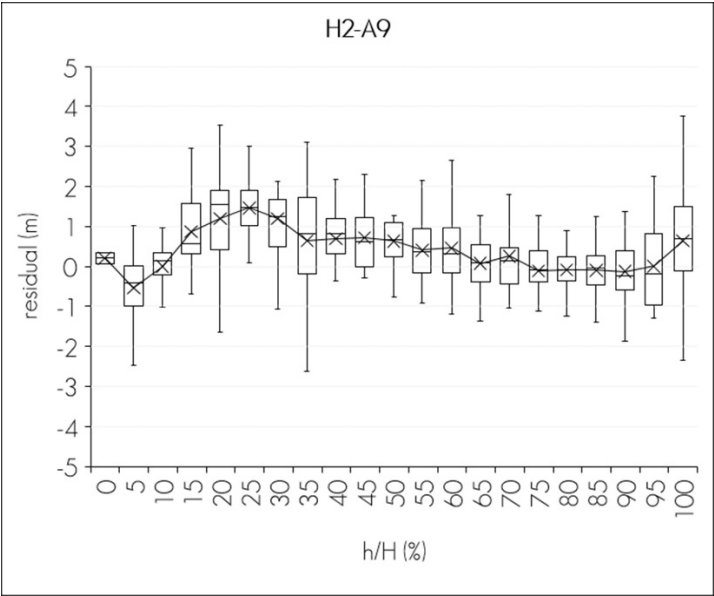


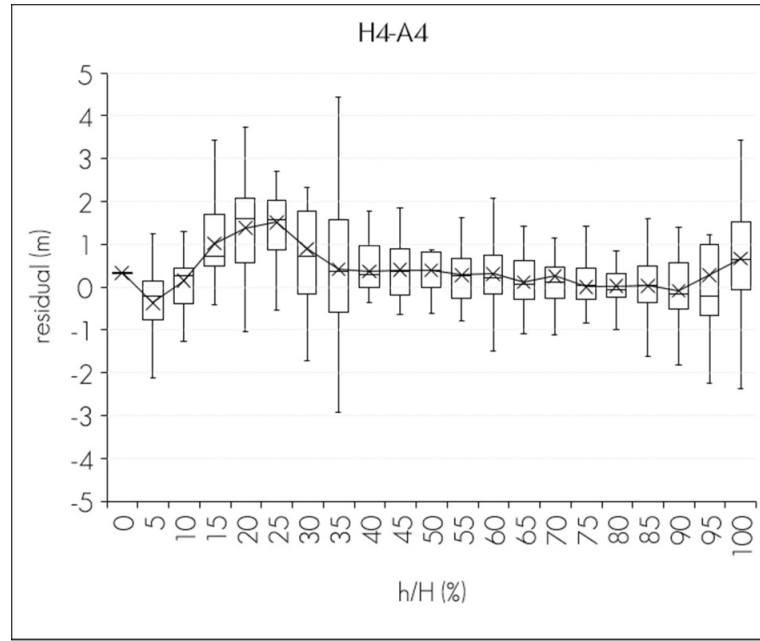
H1 = varPower; H2 = varExp; H3 = varConstPower; H4 = varComb;
 A5 = corARMA1-1; A6 = corARMA2-2; A8 = corARMA3-1; A9 = corARMA3-2.

Figure 1. Box and whisker plots for the distribution of the taper residuals by relative height for combinations of varFunc and corStruct.

In dominant height growth and site index equations with an algebraic difference approach (ADA) or generalized (GADA) equations, the power or exponential functions have been used for the correlation of the heteroscedasticity, in which unequal variances are assumed to exist in the generalized least squares adjustment process (Castillo *et al.*, 2013; Quiñonez-Barraza *et al.*, 2015; Rodríguez-Carrillo *et al.*, 2015; González *et al.*, 2016), in addition to continuous-time autoregressive structures of the errors, in order to correct the autocorrelation with tree stem analysis of various species in natural forests and commercial forest plantations, such as *Pinus durangensis* Martínez, *P. arizonica* Engelm., *Juniperus deppeana* Steud, *P. teocote* Schiede ex Schltdl. & Cham. and *P. ayacahuite* Ehrenb. ex Schltdl. The approach developed in this paper considers the correction of the heteroscedasticity and of the autocorrelation as combinations of functions (Rodríguez *et al.*, 2013; Williams *et al.*, 2013).







H1 = varPower; H2 = varExp; H3 = varConstPower; H4 = varComb;
A4 = corARMA2-0; A6 = corARMA2-2; A7 = corARMA2-2; A9 = corARMA3-2.

Figure 2. Box and whisker plots for the distribution of the residuals of height growth by relative height for combinations of varFunc and corStruct.

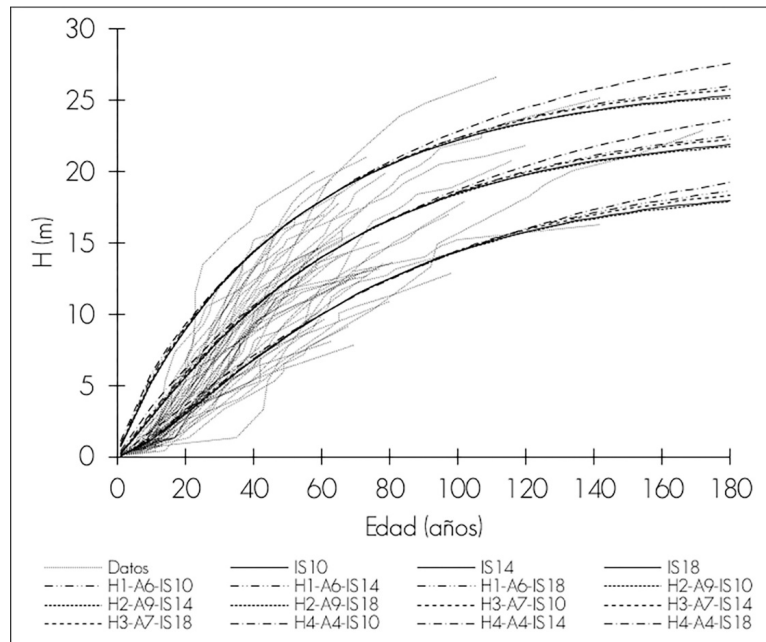
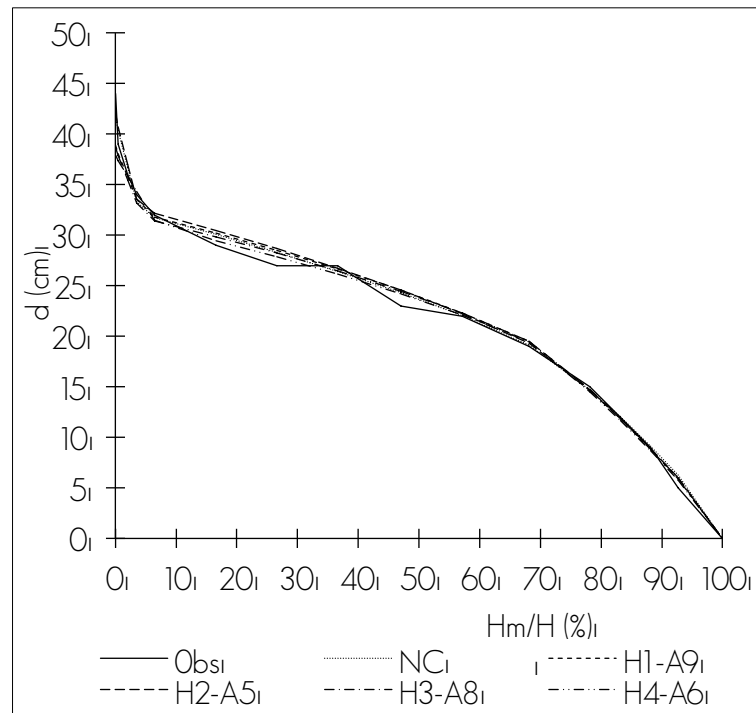
Table 4. Parameter estimates of the top four combinations of the variance functions and autocorrelation structures in the taper (T) and height growth (HG) equations.

Combination	St	a_0	a_1	a_2	β_1	β_2	β_3	p_1	p_2	ρ_1	ρ_2	ρ_3	θ_1	θ_2	δ_1	δ_2
A-H1-A9	φ	7.4×10^{-5}	1.863	0.968	8.4×10^{-6}	4.1×10^{-5}	2.9×10^{-5}	0.045	0.708	0.944	0.784	-0.760	-0.338	-0.649	-0.129	
	ω	9.2×10^{-5}	0.057	0.069	5.5×10^{-7}	6.4×10^{-7}	4.8×10^{-7}	0.002	0.008	0.013	0.057	0.070	0.746	0.142	0.006	
A-H2-A5*	φ	8.8×10^{-5}	1.704	1.099	5.8×10^{-6}	4.0×10^{-5}	2.9×10^{-5}	0.040	0.709	0.725	-0.052		-0.025		-0.024	
	ω	1.1×10^{-5}	0.060	0.745	2.4×10^{-7}	5.7×10^{-7}	5.4×10^{-7}	0.001	0.010	0.039	0.039		3.23		0.002	
A-H3-A8	φ	7.4×10^{-5}	1.873	0.953	8.2×10^{-6}	4.1×10^{-5}	2.9×10^{-5}	0.045	0.708	0.586	0.540	-0.351	0.026		2.1×10^{-8}	-0.128
	ω	9.9×10^{-6}	0.060	0.073	5.2×10^{-7}	6.1×10^{-7}	4.9×10^{-7}	0.002	0.007	0.774	0.738	0.116	0.406		1.8×10^{-8}	0.006
A-H4-A6	φ	7.5×10^{-5}	1.935	0.871	1.2×10^{-5}	4.2×10^{-5}	2.9×10^{-5}	0.072	0.701	-0.176	0.785		0.880		0.045	-0.191
	ω	8.7×10^{-6}	0.052	0.063	5.6×10^{-7}	5.5×10^{-7}	5.9×10^{-7}	0.002	0.007	0.097	0.077		0.745		0.005	0.010
CA-H1-A6*	φ				3.738	-0.523	0.015			0.974	-0.005		-0.528		-0.087	
	ω				0.141	0.122	0.001			0.200	0.306		0.139		0.027	

CA-H2-A9*	φ	3.702	-0.488	0.018	0.117	0.921	-0.087	-0.316	-0.521	-0.033	
	ω	0.121	0.100	0.001	0.441	0.563	-0.158	0.258	0.095	0.004	
CA-H3-A7*	φ	3.739	-0.522	0.016	0.006	0.929		0.404	-0.453	1.2x10 ⁻⁶	-0.080
	ω	0.141	0.122	0.001	0.066	0.057		0.154	0.062	6.3x10 ⁻⁶	0.027
CA-H4-A4*	φ	3.994	-0.769	0.011	0.291	0.647				-0.114	0.477
	ω	0.156	0.183	0.001	0.031	0.056				0.010	0.055

St = Statistic; φ = Parameter; ω = Parameter estimator; ρ_i = Parameters of the continuous autocorrelation structures (corCompSymm, AR y CAR) ($i=1, 2, 3$); θ_1 = Parameters of the moving average (MA) autocorrelation structures (MA) ($i=1, 2, 3$); δ_1 = Parameters of the power functions ($i= 1, 2$); * = Combinations with constant variances according to Bartlett's test of homogeneity of variances. H1 = varPower; H2 = varExp; H3 = varConstPower; H4 = varComb; A4 = corARMA2-0; A5 = corARMA1-1; A6 = corARMA2-1; A7 = corARMA2-2; A8 = corARMA3-1; A9 = corARMA3-2.

Figure 3 contrasts the predictions of taper and height growth equations with the selected combinations of power variance functions and correlation structures; it shows the observed tendency of the profile of a tree, the fitted equation without heteroscedasticity and autocorrelation correction (NC) and the combinations H1-A9, H2-A5, H3-A8 and H4A6. Only the combination H2-A5 exhibited constant variances. With regard to the height growth equation, the comparison between three SI curves for the fitted equation without correction (*IS10*, *IS14* and *IS18*) and also the equations with corresponding combinations H1-A6, H2-A9, H3-A7 and H4-A4 are graphically represented. Furthermore, this figure shows the growth tendencies of the trees in dataset; in this case, all the combinations displayed constant variances, whereby the desirable properties in the estimated parameters are guaranteed (Beale *et al.*, 2010; Vogelsang, 2012).



$H1 = \text{varPower}$; $H2 = \text{varExp}$; $H3 = \text{varConstPower}$; $H4 = \text{varComb}$; $A4 = \text{corARMA2-0}$; $A5 = \text{corARMA1-1}$; $A6 = \text{corARMA2-1}$; $A7 = \text{corARMA2-2}$; $A8 = \text{corARMA3-1}$; $A9 = \text{corARMA3-2}$.

Figure 3. Taper prediction charts for a tree profile, and height growth curves by site index (SIs 10, 14 and 18 m at the baseline age of 60 years) with the combinations of variance functions and autocorrelation structures, and the equations without variance or autocorrelation (IS_{10} , IS_{14} , IS_{18}) structures.

Conclusions

The variance functions in combination with the correlation structures corrected the heteroscedasticity assumptions of variances and error autocorrelation in the taper and dominant height growth equations, using a generalized nonlinear least square approach; as a result, unbiased parameters with a minimum variance were obtained.

The predictions of the selected taper equations are more efficient, with consistent fitting statistics; the combination of the exponential variance function with an autoregressive-moving average correlation structure (corARMA1-1) produces constant variances by relative height categories of the stem profiles. The dominant height and site index model, at a base age of 60 years, exhibits realistic predictions. The selected combinations (H1-A6, H2-A9, H3-A7 and H4-A4) exhibit constant variances by relative height categories at specific ages. In both, the residuals are independent, and the properties of the tests of hypothesis of the estimated parameters are guaranteed.

The use of compatible taper and commercial volume equations, as well as of height growth and index site equations, is defined by the use of the intrinsic parameters in each equation; therefore, the parameters of the variance functions and correlation structures are only statistical indicators for rendering the fitting more efficient.

Acknowledgements

The authors wish to express their gratitude to *San Diego de Tezains ejido, Santiago Papasquiaro, Durango, Mexico*, for making the taper and height growth information available to be used in this study.

Conflict of interests

The authors declare no conflict of interests.

Contribution by author

Gerónimo Quiñonez-Barraza: data analysis, fitting of equation, drafting and review of the manuscript; Guadalupe Geraldine García-Espinoza: data analysis, fitting of equation, drafting and review of the manuscript; Oscar Alberto Aguirre-Calderón: data analysis, fitting of equation, drafting and review of the manuscript.

References

- Aguirre-Calderón, O. A. 2015. Manejo Forestal en el Siglo XXI. *Madera y Bosques* 21(21):17-28.
- Arias-Rodil, M., U. Diéguez-Aranda, F. Rodríguez P., C. A. López-Sánchez, E. Canga L., A. Cámara O. and F. Castedo-Dorado. 2015. Modelling and localizing a stem taper function for *Pinus radiata* in Spain. *Canadian Journal of Forest Research* 45(6):647-658.
- Bartlett, M. S. 1954. A Note on the Multiplying Factors for Various χ^2 Approximations. *Journal of the Royal Statistical Society. Series B (Methodological)* 16(2):296-298.
- Beale, C. M., J. J. Lennon, J. M. Yearsley, M. J. Brewer and D. A. Elston. 2010. Regression analysis of spatial data. *Ecology Letters* 13(2):246-264.
- Castillo L., A., B. Vargas-Larreta, J. J. Corral R., J. A. Nájera L., F. Cruz C. y F. J. Hernández. 2013. Modelo compatible altura-índice de sitio para cuatro especies de

pino en Santiago Papasquiaro, Durango. Revista Mexicana de Ciencias Forestales 4(18):89-103.

Corral-Rivas, J. J., D. J. Vega-Nieva, R. Rodríguez-Soalleiro, C. A. López-Sánchez, C. Wehenkel, B. Vargas-Larreta, J. G. Álvarez-González and A. Ruiz-González. 2017. Compatible system for predicting total and merchantable stem volume over and under bark, branch volume and whole-tree volume of Pine species. Forests 8(11):417.

Crecente-Campo, F., J. G. Álvarez-González, F. Castedo-Dorado, E. Gómez-García and U. Diéguez-Aranda. 2013. Development of crown profile models for *Pinus pinaster* Ait. and *Pinus sylvestris* L. in northwestern Spain. Forestry: An International Journal of Forest Research 86(4):481-491.

Durbin, J. and G. S. Watson. 1971. Testing for serial correlation in least squares regression. III. Biometrika 58(1):1-19.

Fang, Z., B. E. Borders and R. L. Bailey. 2000. Compatible volume-taper models for Loblolly and Slash pine based on a system with segmented-stem form factors. Forest Science 46(1):1-12.

Fierros-Mateo, R., H. M. De los Santos-Posadas, M. A. Fierros-González y F. Cruz-Cobos. 2017. Crecimiento y rendimiento maderable en plantaciones de *Pinus chiapensis* (Martínez) Andresen. Agrociencia 51(2):201-214.

Fortin, M., G. Daigle, C.-H. Ung, J. Bégin and L. Archambault. 2007. A variance-covariance structure to take into account repeated measurements and heteroscedasticity in growth modeling. European Journal of Forest Research 126(4):573-585.

García, E. M. 1981. Modificaciones al Sistema de Clasificación Climática de Köppen. México, D.F., México. 500 p.

González M., M., F. Cruz C., G. Quiñonez B., B. Vargas L. y J. A. Nájera L. 2016. Modelo de crecimiento en altura dominante para *Pinus pseudostrobus* Lindl. en el estado de Guerrero. Revista Mexicana de Ciencias Forestales 7(37):7-20.

- Gregoire, T. G. and O. Schabenberger. 1996. A non-linear mixed-effects model to predict cumulative bole volume of standing trees. *Journal of Applied Statistics* 23(2-3):257-272.
- Gujarati, D. N. y D. C. Porter. 2011. *Econometria Básica*. AMGH Editora. México, D.F., México. pp. 365, 371, 388 y 413.
- Hidding, B., J. P. Tremblay and S. D. Côté. 2013. A large herbivore triggers alternative successional trajectories in the boreal forest. *Ecology* 94(12):2852-2860.
- Krisnawati, H. 2016. A compatible estimation model of stem volume and taper for *Acacia mangium* Willd. plantations. *Indonesian Journal of Forestry Research* 3(1):49-64.
- Layard, M. 1973. Robust large-sample tests for homogeneity of variances. *Journal of the American Statistical Association* 68(341):195-198.
- Özçelik, R. and F. Crecente-Campo. 2016. Stem taper equations for estimating merchantable volume of Lebanon Cedar trees in the Taurus Mountains, Southern Turkey. *Forest Science* 62(1):78-91.
- Paulo, J. A., J. H. N. Palma, A. A. Gomes, S. P. Faias, J. Tomé and M. Tomé. 2015. Predicting site index from climate and soil variables for cork oak (*Quercus suber* L.) stands in Portugal. *New Forests* 46(2):293-307.
- Pinheiro, J. and D. Bates. 2000. *Mixed-effects models in S and S-PLUS*. Statistics and computing. Springer-Verlag. New York, NY USA. 528 p.
- Pinheiro, J., D. Bates, S. DebRoy, D. Sarkar and R. C. Team. 2015. nlme: Linear and Nonlinear Mixed Effects Models. R package version 3.1-120. <https://CRAN.R-project.org/package=nlme> (4 de diciembre de 2017).
- Quiñonez-Barraza, G., H. M. De los Santos-Posadas, J. G. Álvarez-González y A. Velázquez-Martínez. 2014. Sistema compatible de ahusamiento y volumen comercial para las principales especies de *Pinus* en Durango, México. *Agrociencia* 48(5):553-567.
- Quiñonez-Barraza, G., H. M. De los Santos-Posadas, F. Cruz-Cobos, A. Velázquez-Martínez, G. Ángeles-Pérez y G. Ramírez-Valverde. 2015. Índice de sitio con

polimorfismo complejo para masas forestales de Durango, México. *Agrociencia* 49(4):439-454.

R Core Team. 2017. R: A language and environment for statistical computing. Editorial. Vienna, Austria. 3475 p.

Rodríguez-Carrillo, A., F. Cruz-Cobos, B. Vargas-Larreta and F. J. Hernández. 2015. Compatible dominant height-site index model for juniper (*Juniperus deppeana* Steud.). *Revista Chapingo. Serie Ciencias Forestales y del Ambiente* 21(1):97-108.

Rodríguez, F., I. Lizarralde and F. Bravo. 2013. Additivity on nonlinear stem taper functions: A case for corsican Pine in Northern Spain. *Forest Science* 59(4):464-471.

Sakici, O. E., N. Misir, H. Yavuz and M. Misir. 2008. Stem taper functions for *Abies nordmanniana* subsp. *bornmulleriana* in Turkey. *Scandinavian Journal of Forest Research* 23(6):522-533.

Salas, C., T. G. Gregoire, D. J. Craven y H. Gilabert. 2016. Modelación del crecimiento de bosques: estado del arte. *Bosque* 37(1):3-12.

Sharma, M., N. Subedi, M. Ter-Mikaelian and J. Parton. 2015. Modeling climatic effects on stand height/site index of Plantation-Grown Jack Pine and Black Spruce trees. *Forest Science* 61(1):25-34.

Tamarit U., J. C., E. Rojas D., G. Quiñonez B., C. Ordoñez P. y J. C. Monárrez G. 2017. Sistema de cubicación para árboles individuales de *Quercus* sp. en bosques bajo manejo de Puebla, México. *Revista Mexicana de Ciencias Forestales* 8(40):69-88.

Tamarit U., J. C., H. M. De los Santos P., A. Aldrete, J. R. Valdez L., H. Ramírez M. y V. Guerra D. 2014. Sistema de cubicación para árboles individuales de *Tectona grandis* L. f. mediante funciones compatibles de ahusamiento-volumen. *Revista Mexicana de Ciencias Forestales* 5(21):58-74.

Tang, X., C. Pérez-Cruzado, L. Fehrmann, J. G. Álvarez-González, Y. Lu and C. Kleinn. 2016. Development of a compatible taper function and stand-level merchantable volume model for Chinese fir plantations. *PloS one* 11(1):e0147610.

Uranga-Valencia, L. P., H. M. De los Santos-Posadas, J. R. Valdez-Lazalde, J. López-Upton y H. Navarro-Garza. 2015. Volumen total y ahusamiento para *Pinus patula* Schiede ex Schltdl. et Cham. en tres condiciones de bosque. *Agrociencia* 49(7):787-801.

Vogelsang, T. J. 2012. Heteroskedasticity, autocorrelation, and spatial correlation robust inference in linear panel models with fixed-effects. *Journal of Econometrics* 166(2):303-319.

Williams, M. N., C. A. Gómez G. and D. Kurkiewicz. 2013. Assumptions of multiple regression: correcting two misconceptions. *Practical Research & Evaluation* 18(11):1-14.

Xu, H., Y. Sun, X. Wang, Y. Fu, Y. Dong and Y. Li. 2014. Nonlinear mixed-effects (NLME) diameter growth models for individual China-fir (*Cunninghamia lanceolata*) trees in Southeast China. *PloS one* 9(8):e104012.